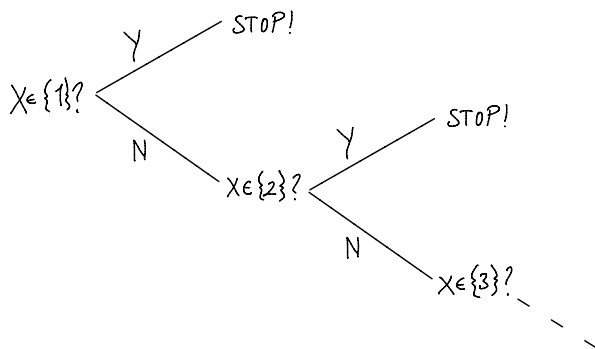


2.1
(a) $X \sim \text{Geom}(1/2)$ $P_X(k) = \left(\frac{1}{2}\right)^k$, $E[X] = 2$

(b)



1. Set $i=1$
2. Ask " $\exists x \in \{i\}?$ ". Wait for answer.
3. If the answer is "No" increment i and go to step 2.
If the answer is "Yes" STOP.

$$E[\# \text{ of Y/N questions}] = \sum_{k=1}^{\infty} k \left(\frac{1}{2}\right)^k = E[X] = 2 = H(X) \quad (*)$$

Note: \otimes suggests that the above scheme of Y/N questions is optimal.

2.4

- (a) By chain rule for entropy.
 (b) $g(X)$ is a deterministic function of X . Thus, given X , there is nothing random regarding $g(X) \Rightarrow H(g(X)|X)=0$
 (c) By chain rule of entropy.
 (d) Conditional entropy is non-negative. Equality holds iff g^{-1} is well-defined (i.e., when g is bijection)

2.8

→ Drawing $k \geq 2$ balls with replacement has higher entropy here is why:

Let X_i denote the random variable of drawing a ball from the urn the i -th time w/ replacement, and Y_i denote the random variable of drawing a ball from the urn the i -th time w/out replacement.

Note that X_1, X_2, \dots, X_k are i.i.d random variables, whereas Y_i depends on Y^{i-1} $\forall i \geq 2$. Note further that X_i and Y_i are identically distributed

$$H(X_1, \dots, X_k) = \sum_{i=1}^k H(X_i) \quad \text{and} \quad H(Y_1, \dots, Y_k) = \sum_{i=1}^k H(Y_i | Y^{i-1})$$

Since conditioning reduces entropy, we have $H(Y_i) \geq H(Y_i | Y^{i-1})$ $\forall i \geq 2$ and since Y_i and X_i are identically distributed it follows that $H(X_i) = H(Y_i) \geq H(Y_i | Y^{i-1})$ ■

2.9

(a) Let $\rho(X, Y) = H(X|Y) + H(Y|X)$

(i) $\rho(X, Y) \geq 0$ as conditional entropy is non-negative

(ii) $\rho(X, Y) = \rho(Y, X)$ is clear from the definition

(iii) If equality $X=Y$ means $\exists f$ a bijection such that $f(X)=Y$, then

• when $X=Y$; $H(X|f(X)) + H(f(X)|X) = 0$ and

• $\rho(X, Y) = 0 \Rightarrow H(X|Y) = 0$ and $H(Y|X) = 0 \Rightarrow \exists f$, a bijection such that $f(X)=Y$.

$$(iv) \quad H(X|Y) + H(Y|X) + H(Y|Z) + H(Z|Y) \geq H(X, Y|Z) + H(X, Z|X) \quad (*)$$

$$\geq H(X|Z) + H(Z|X) \quad (**)$$

where (*) follows from the followings

$$H(X|Y) \geq H(X|Y, Z)$$

$$H(Z|Y) \geq H(Z|X, Y)$$

and (**) follows from the fact that joint entropy is no-less than marginal entropies.
(in other words, (**) follows because $H(Y|X, Z) \geq 0$)

(b) Note that $I(X; Y) \stackrel{(1)}{=} H(X) - H(X|Y) \stackrel{(2)}{=} H(Y) - H(Y|X) \stackrel{(3)}{=} H(X) + H(Y) - H(X, Y)$.

• So $H(X) + H(Y) - 2I(X; Y) = H(X|Y) + H(Y|X) = \rho(X, Y)$ Hence, (2.172) follows from (1) and (2)

Note further that $H(X, Y) \stackrel{(4)}{=} H(X) + H(Y) - I(X; Y)$.

• Using (2.172), and (4), (2.173) is now verified.

• (2.174) follows from (3) and (2.173)

2.10 (a) Let $f(X) = \mathbb{1}\{X = X_1\}$

Note that $H(X) = H(X, f(X)) = H(f(X)) + H(X|f(X))$

$$= H(\alpha) + \alpha H(X_1) + (1-\alpha) H(X_2)$$

where $H(\alpha)$ denotes the binary entropy function.

(b)

$$\frac{d}{d\alpha} H(\alpha) = 0 \Rightarrow \log_2 \frac{1-\alpha}{\alpha} + H(X_1) - H(X_2) = 0 \Rightarrow \alpha = \frac{1}{1 + 2^{-(H(X_1) - H(X_2))}}$$

Note that $H(X)$ is a concave function of α . (because $H(\alpha)$ is concave and $\alpha H(X_1) + (1-\alpha) H(X_2)$ is linear wrt. α)
As a result, global maximum is achieved at $\alpha = \frac{1}{1 + 2^{-(H(X_1) - H(X_2))}}$

Let $\xi = H(X_1) - H(X_2)$ and $\alpha^* = \frac{1}{1+2^{-\xi}}$ and note that

$$\begin{aligned} H(X) &\leq H(\alpha^*) + \alpha^* H(X_1) + (1-\alpha^*) H(X_2) \\ &= \log_2(1+2^{-\xi}) + (1-\alpha^*) \xi + \alpha^* H(X_1) + (1-\alpha^*) H(X_2) \\ &= \log_2(1+2^{-H(X_1)+H(X_2)}) + H(X_1) = \log_2(2^{H(X_1)} + 2^{H(X_2)}) \\ \Rightarrow 2^{H(X)-H(X_1)} &\leq 1+2^{-H(X_1)+H(X_2)} \Rightarrow 2^{H(X)} \leq 2^{H(X_1)} + 2^{H(X_2)} \end{aligned}$$

(Note that since $\frac{d}{d\alpha} H(\alpha) = \log \frac{1-\alpha}{\alpha}$ for any arbitrary logarithm base, the following result comes for free)

$$b^{H(X)} \leq b^{H(X_1)} + b^{H(X_2)} \quad \forall b > 1$$

2.17

(a) X_i are i.i.d. $\text{Ber}(p)$

(b) $H(X_1, \dots, X_n) \geq H(f(X_1, \dots, X_n))$ where $f(X_1, \dots, X_n) = (Z_1, \dots, Z_K)$ with random K

(c) Chain rule: $H(V, W) = H(W) + H(V|W)$

(d) Z_i are i.i.d. $\text{Ber}(1/2)$ $H(Z_1, \dots, Z_K | K=k) = k$ bits so $H(Z_1, \dots, Z_K | K) = E_K[H(Z_1, \dots, Z_K | K)] = E[K]$

(e) $H(K) \geq 0$

A good map f on sequences of length 4:

• To maximize the expected number of pure bits, we combine non-pure bits of same probability. Note that in (X_1, \dots, X_4)

$$\begin{array}{lcl} 0000 \rightarrow \text{occurs with probability } (1-p)^4 & 0001 \\ 1111 \rightarrow \text{occurs with probability } p^4 & 0010 \\ 0011 \left. \begin{array}{l} 0101 \\ 1001 \\ 0110 \\ 1010 \\ 1100 \end{array} \right\} \rightarrow \text{occurs with probability } p^2(1-p)^2 & \left. \begin{array}{l} 0100 \\ 1000 \\ 1110 \\ 1101 \\ 1011 \\ 0111 \end{array} \right\} \rightarrow \text{occurs with probability } p(1-p)^3 \\ & & \left. \begin{array}{l} 1101 \\ 1011 \\ 0111 \end{array} \right\} \rightarrow \text{occurs with probability } p^3(1-p) \end{array}$$

Therefore, let f be the following map:

$$\begin{array}{llllll} 0000 \rightarrow \wedge & 0001 \left. \begin{array}{l} 1110 \\ 1100 \end{array} \right\} \rightarrow 00 & 0010 \left. \begin{array}{l} 1101 \\ 0101 \end{array} \right\} \rightarrow 01 & 0100 \left. \begin{array}{l} 1011 \\ 1010 \end{array} \right\} \rightarrow 10 & 1000 \left. \begin{array}{l} 0111 \\ 0011 \end{array} \right\} \rightarrow 11 & 1001 \rightarrow 1 \\ 1111 \rightarrow \wedge & & & & & 0110 \rightarrow 0 \end{array}$$

This should be a good map as we utilize every possible outcome of (X_1, \dots, X_4) except that 0000 and 1111 which are the most unlikely/likely sequence (depending on the value of p).

2.37 $D(P_{XYZ} \| P_X P_Y P_Z) = H(X) + H(Y) + H(Z) - H(X, Y, Z)$

The given quantity is equal to 0 iff $P_{XYZ} = P_X P_Y P_Z$, i.e., $D(P_{XYZ} \| P_X P_Y P_Z) = 0$ iff X, Y and Z are independent from each other.

$$D(P_{XYZ} \| P_X P_Y P_Z) = E \left[\log \frac{P_{XYZ}(X, Y, Z)}{P_X(X) P_Y(Y) P_Z(Z)} \right] = E \left[\log \frac{P_{XY}(X, Y)}{P_X(X) P_Y(Y)} + \log \frac{P_{Z|XY}(Z|X, Y) P_{XY}(X, Y)}{P_Z(Z) P_{XY}(X, Y)} \right]$$

$$= I(X; Y) + I(Z; X, Y) \quad (*)$$

[or using similar derivation (*) is equal to: $I(X; Z) + I(Y; X, Z)$ and $I(Y; Z) + I(X; Y, Z)$]

2.41

(a) $I(X; Q, A) = I(X; Q) + I(X; A|Q)$ (chain rule for mutual information)

$$= I(X; A|Q) \quad (X \perp\!\!\!\perp Q)$$

$$= H(A|Q) - H(A|Q, X)$$

$$= H(A|Q) \quad (H(A|Q, X) = 0 \text{ as } A \text{ is deterministic function of } Q \text{ and } X.)$$

Interpretation: The uncertainty removed by (Q, A) pair is the same as the average information of the answer A given the question Q . (Note that information of an outcome x of a discrete r.v. X is given by $I_X(x) = \log \frac{1}{P_X(x)}$.)

(b) $I(X; Q_1, A_1, Q_2, A_2) = I(X; Q_1, A_1) + I(X; Q_2, A_2|Q_1, A_1)$

So it suffices to show that $I(X; Q_2, A_2|Q_1, A_1) \leq H(A_2|Q_2) = I(X; Q_2, A_2)$

In that respect, note that

$$I(X; Q_2, A_2|Q_1, A_1) = I(X; Q_2|Q_1, A_1) + I(X; A_2|Q_1, A_1, Q_2) \quad (1)$$

$$= I(X; A_2|Q_1, A_1, Q_2) \quad (2)$$

$$= H(A_2|Q_1, A_1, Q_2) - H(A_2|Q_1, A_1, Q_2, X) \quad (3)$$

$$= H(A_2|Q_1, A_1, Q_2) \quad (4)$$

$$\leq H(A_2|Q_2) \quad (5)$$

where (1) follows from chain rule, (2) follows because X does not depend on Q_2 , (4) follows because given Q_2 and X , A_2 is deterministic, (5) is due to the fact that further conditioning cannot increase entropy.

3.4 X_i iid $\sim p(x)$ on $\{1, 2, \dots, m\}$, $\mu = E[X]$, $H = \sum p(x) \log \frac{1}{p(x)}$
 $A_n = \{x^n \in X^n : |\frac{1}{n} \log \frac{1}{p(x^n)} - H| \leq \epsilon\}$ and $B^n = \{x^n \in X^n : |\frac{1}{n} \sum_{i=1}^n x_i - \mu| \leq \epsilon\}$

- (a) $P[X^n \in A^n] \rightarrow 1$ as $n \rightarrow \infty$, because of AEP Theorem
 (b) $P[X^n \in A^n \cap B^n] = 1 - P[X^n \in (A^n)^c \cup B^n]^c \geq 1 - (P[X^n \in (A^n)^c] + P[X^n \in B^n]^c)$
 $P[X^n \in (A^n)^c] \rightarrow 0$ by AEP Theorem, $P[X^n \in B^n]^c \rightarrow 0$ by LLN
 It follows that $P[X^n \in A^n \cap B^n] \rightarrow 1$.
 (c) $1 \geq \sum_{x^n \in A^n \cap B^n} p(x^n) \geq |A^n \cap B^n| \cdot 2^{-n(H+\epsilon)} \Rightarrow |A^n \cap B^n| \leq 2^{n(H+\epsilon)} \quad \forall n$
 \hookrightarrow by definition of A^n .

(d) Since $P[X^n \in A^n \cap B^n] \rightarrow 1$, for sufficiently large n , we have
 $\frac{1}{2} \leq P[X^n \in A^n \cap B^n] = \sum_{x^n \in A^n \cap B^n} p(x^n) \leq |A^n \cap B^n| 2^{-n(H-\epsilon)} \Rightarrow |A^n \cap B^n| \geq \frac{1}{2} 2^{n(H-\epsilon)}$ for sufficiently large n .
 \hookrightarrow by definition of A^n

3.10

$$V_n^{1/n} = \left(\prod_{i=1}^n X_i \right)^{1/n} \Rightarrow \frac{1}{n} \log V_n = \frac{1}{n} \sum_{i=1}^n \log X_i \xrightarrow[n \rightarrow \infty]{LLN} E[\log X_1] = \int_0^1 \log(x) dx = -1$$

$$\lim_{n \rightarrow \infty} V_n^{1/n} = \lim_{n \rightarrow \infty} e^{\log V_n^{1/n}} = e^{-1} \quad (\text{because } e^x \text{ is a continuous function})$$

$$E[V_n] = \int_0^1 \dots \int_0^1 (x_1 x_2 \dots x_n) dx_1 dx_2 \dots dx_n \quad E[V_n] = \left(\frac{1}{2}\right)^n \quad \lim_{n \rightarrow \infty} (E[V_n])^{1/n} = 1/2 > 1/e$$

3.11

(a) $P[A \cap B] = 1 - P[A^c \cup B^c] \geq 1 - P[A^c] - P[B^c] \geq 1 - \epsilon_1 - \epsilon_2$
 \hookrightarrow union bound $\hookrightarrow P[A^c] \leq \epsilon_1$ and $P[B^c] \leq \epsilon_2$

(b) $1 - \epsilon - \delta \leq P[A_\epsilon^{(n)} \cap B_\delta^{(n)}]$ (follows from part (a))
 $= \sum_{A_\epsilon^{(n)} \cap B_\delta^{(n)}} p(x^n)$ (definition of $P[A_\epsilon^{(n)} \cap B_\delta^{(n)}]$)
 $\leq \sum_{A_\epsilon^{(n)} \cap B_\delta^{(n)}} 2^{-n(H-\epsilon)}$ (follows from the definition of the typical set $A_\epsilon^{(n)}$)
 $= |A_\epsilon^{(n)} \cap B_\delta^{(n)}| 2^{-n(H-\epsilon)}$ (term inside summation no longer depends on summation index x^n)
 $\leq |B_\delta^{(n)}| 2^{-n(H-\epsilon)}$ ($A_\epsilon^{(n)} \cap B_\delta^{(n)} \subseteq B_\delta^{(n)}$)

(c) By part (b), we have $|B_\delta^{(n)}| \geq 2^{n(H-\epsilon)} (1 - \epsilon - \delta)$, That is for sufficiently large n , $|B_\delta^{(n)}| > 2^{n(H-\epsilon)}$ which is the promised result of Theorem 3.3.1.

3.13

(a) $H(X) = 0.970951$ bits.

(b) $x^{25} \in A_{0.1}^{(25)}$ iff $0.870951 \leq \frac{1}{25} \log \frac{1}{p(x^{25})} \leq 1.070951$ (*)

By looking at the given table we see that sequences which have more than 11 one's and less than 19 one's satisfy (*). (This part of the table was correct)

There are $\binom{25}{11} + \binom{25}{12} + \binom{25}{13} + \binom{25}{14} + \binom{25}{15} + \binom{25}{16} + \binom{25}{17} + \binom{25}{18} + \binom{25}{19} = 26366510$ elements in $A_{0.1}^{(25)}$.

Probability of the typical set is $= 0.936266$

(c) Smallest set that has probability 0.9 contains elements with $k \geq 13$ one's and 3680638 elements with exactly 12 one's.

So there are 20457854 elements in that set.

(d) The intersection contains 20389448 elements.

The probability of the intersection is roughly ≈ 0.870638